

Exploratory Factor Analysis

```
library(tidyverse)
library(psych)
library(dagitty)
library(plotly)
library(lavaan)
library(EFAtools)
library(flextable)
```

Path Diagram

Personally, it may be easier to use powerpoint or other tools with a graphical interface. But if you want to do it with code, here is an example with [Graphviz](#).

Graphviz

```
digraph FactorModel {
  node [shape=circle,fontsize=12];
  F1; F2;

  node [shape=rectangle];
  x1; x2; x3; x4; x5; x6;

  edge [fontsize=10];
  F2 -> F1 [label="r" constraint=false dir="both"];
  F1 -> x1 [label="a"];
  F1 -> x2 [label="b"];
  F1 -> x3 [label="c"];
  F1 -> x4 [label="d"]
```

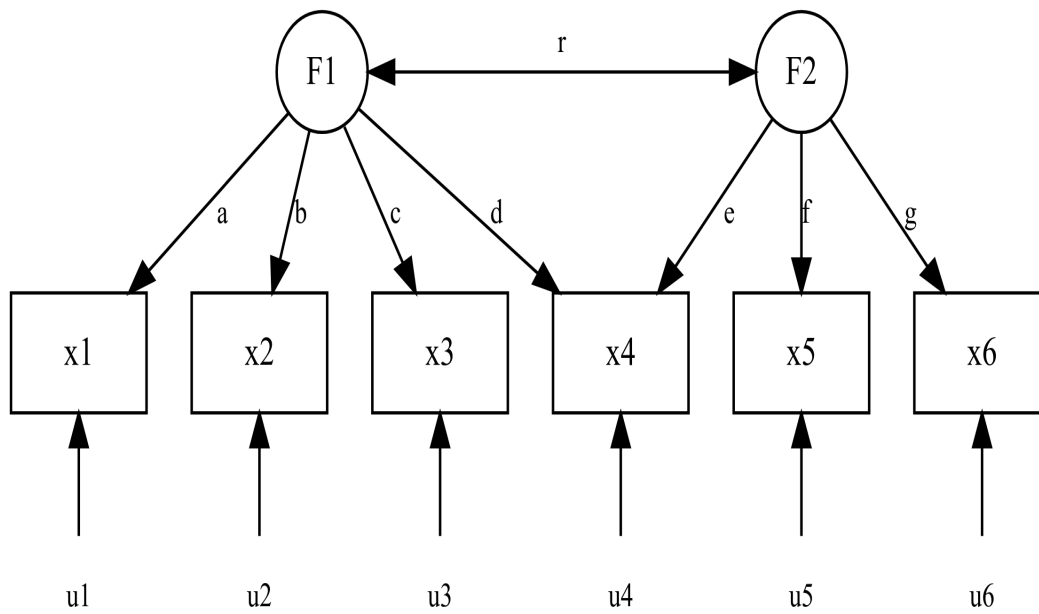
```

F2 -> x4 [label="e"];
F2 -> x5 [label="f"];
F2 -> x6 [label="g"];

node [shape=none,fontsize=10];
u1 -> x1;
u2 -> x2;
u3 -> x3;
u4 -> x4;
u5 -> x5;
u6 -> x6;

{rank=same; x1 x2 x3 x4 x5 x6}
{rank=min; F1 F2}
{rank=max; u1 u2 u3 u4 u5 u6}
}

```

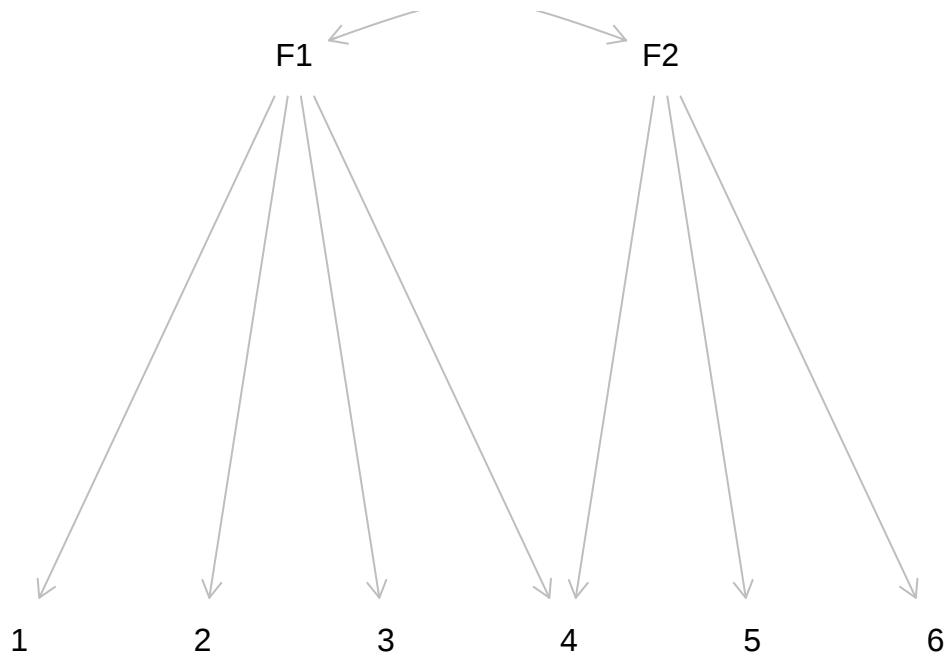


Path diagram of a two-factor model.

DAG

A factor model can also be represented as a directed acyclic graph (DAG). The `dagitty` package can show the vanishing tetrad assumptions:

```
dag1 <- dagitty("dag{ F1 -> {1 2 3 4}; F2 -> {4 5 6}; F1 <-> F2 }")
latents(dag1) <- c("F1", "F2")
coordinates(dag1) <- list(
  x = c(F1 = 0, F2 = 2, `1` = -1.5, `2` = -0.5, `3` = 0.5, `4` = 1.5,
        `5` = 2.5, `6` = 3.5),
  y = c(F1 = 0, F2 = 0, `1` = 1, `2` = 1, `3` = 1, `4` = 1,
        `5` = 1, `6` = 1)
)
plot(dag1)
```



```
vanishingTetrads(dag1)
```

	[,1]	[,2]	[,3]	[,4]
[1,]	"1"	"3"	"4"	"2"
[2,]	"1"	"2"	"4"	"3"
[3,]	"1"	"2"	"3"	"4"
[4,]	"1"	"3"	"5"	"2"

```

[5,] "1" "2" "5" "3"
[6,] "1" "2" "3" "5"
[7,] "1" "3" "6" "2"
[8,] "1" "2" "6" "3"
[9,] "1" "2" "3" "6"
[10,] "1" "4" "5" "2"
[11,] "1" "4" "6" "2"
[12,] "1" "5" "6" "2"
[13,] "1" "4" "5" "3"
[14,] "1" "4" "6" "3"
[15,] "1" "5" "6" "3"
[16,] "1" "5" "6" "4"
[17,] "2" "4" "5" "3"
[18,] "2" "4" "6" "3"
[19,] "2" "5" "6" "3"
[20,] "2" "5" "6" "4"
[21,] "3" "5" "6" "4"

```

Correlation Matrix

```

# Covariance
cov_oa <- cov(
  select(bfi, 01:05, A1:A5),
  # Listwise deletion
  use = "complete"
)
# Correlation
corr_oa <- cor(
  select(bfi, 01:05, A1:A5),
  # Listwise deletion
  use = "complete"
)
round(corr_oa, digits = 2)

```

	01	02	03	04	05	A1	A2	A3	A4	A5
01	1.00	-0.22	0.39	0.19	-0.24	0.01	0.13	0.15	0.06	0.16
02	-0.22	1.00	-0.28	-0.08	0.33	0.07	0.01	0.00	0.04	-0.01
03	0.39	-0.28	1.00	0.19	-0.32	-0.06	0.16	0.23	0.07	0.24
04	0.19	-0.08	0.19	1.00	-0.18	-0.08	0.08	0.04	-0.04	0.02
05	-0.24	0.33	-0.32	-0.18	1.00	0.11	-0.09	-0.06	0.02	-0.06

A1	0.01	0.07	-0.06	-0.08	0.11	1.00	-0.34	-0.27	-0.15	-0.19
A2	0.13	0.01	0.16	0.08	-0.09	-0.34	1.00	0.49	0.33	0.39
A3	0.15	0.00	0.23	0.04	-0.06	-0.27	0.49	1.00	0.37	0.51
A4	0.06	0.04	0.07	-0.04	0.02	-0.15	0.33	0.37	1.00	0.31
A5	0.16	-0.01	0.24	0.02	-0.06	-0.19	0.39	0.51	0.31	1.00

Polychoric Correlation

The items are technically ordinal, but with six points it is usually less a problem to just use Pearson correlation. We can use the `lavaan` package to get the polychoric correlation.

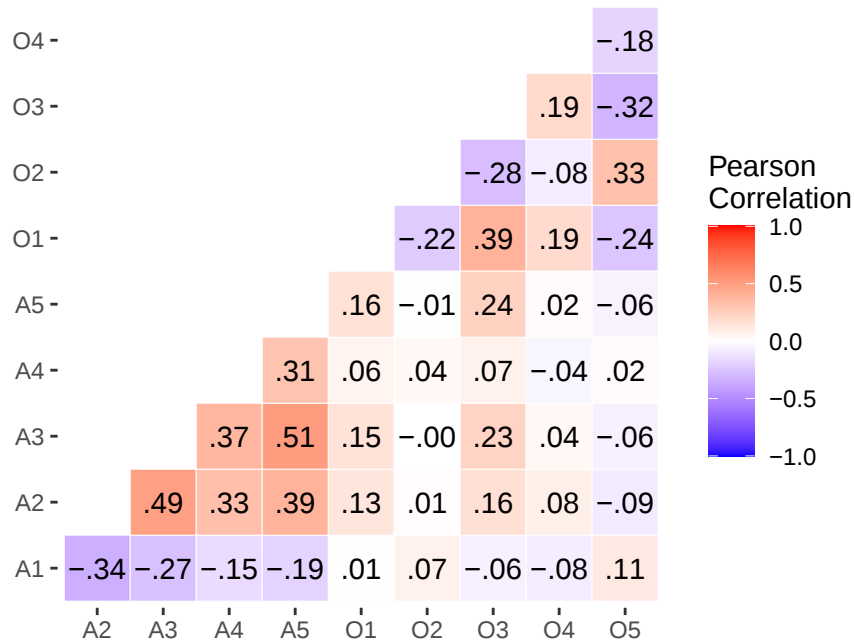
```
pcorr_oa <- lavaan::lavCor(select(bfi, 01:05, A1:A5), ordered = TRUE)
pcorr_oa
```

	01	02	03	04	05	A1	A2	A3	A4	A5
01	1.000									
02	-0.273	1.000								
03	0.450	-0.337	1.000							
04	0.258	-0.106	0.255	1.000						
05	-0.304	0.373	-0.381	-0.250	1.000					
A1	-0.010	0.079	-0.086	-0.100	0.134	1.000				
A2	0.163	-0.002	0.189	0.109	-0.114	-0.411	1.000			
A3	0.187	-0.029	0.270	0.065	-0.089	-0.327	0.565	1.000		
A4	0.071	0.035	0.082	-0.049	0.026	-0.182	0.390	0.422	1.000	
A5	0.195	-0.018	0.275	0.028	-0.074	-0.234	0.453	0.576	0.358	1.000

Heat Map

```
corr_oa |>
  as_tibble(rownames = "Var1") |>
  pivot_longer(cols = -Var1, names_to = "Var2") |>
  mutate(across(1:2, .fns = as.ordered)) |>
  filter(Var2 > Var1) |>
  ggplot(aes(x = Var2, y = Var1, fill = value)) +
  geom_tile(color = "white") +
  geom_text(aes(Var2, Var1, label = rmlpad0(value)), color = "black", size = 4) +
  scale_fill_gradient2(low = "blue", high = "red", limit = c(-1, 1), name = "Pearson\nCorr") +
  theme(axis.title.x = element_blank(),
        axis.title.y = element_blank(),
```

```
panel.background = element_blank()) +
coord_fixed()
```



Factor Extraction

Eigen-Decomposition

A good video to explain this: <https://www.youtube.com/watch?v=86ODrk1nB-g>

```
# Simulate 3-D Data
lambda_mat <- matrix(c(.8, .7, .5, .1, .5, .4), nrow = 3)
sigma_mat <- tcrossprod(lambda_mat)
diag(sigma_mat) <- 1
y <- MASS::mvrnorm(500, mu = rep(0, 3), Sigma = sigma_mat, empirical = TRUE)
```

```
# library(scatterplot3d)
# scatterplot3d(y)
dfy <- as.data.frame(y)
names(dfy) <- c("y1", "y2", "y3")
plot_ly(dfy, x = ~ y1, y = ~ y2, z = ~ y3)
```

No trace type specified:

Based on info supplied, a 'scatter3d' trace seems appropriate.

Read more about this trace type -> <https://plotly.com/r/reference/#scatter3d>

No scatter3d mode specified:

Setting the mode to markers

Read more about this attribute -> <https://plotly.com/r/reference/#scatter-mode>

file:///tmp/RtmpKvOBq0/file32987171e6e980/widget3298716a1d761b.html screenshot completed

Estimating Communalities

Squared multiple correlations (R^2)

Use squared multiple R^2 (not efficient)

```
# E.g., Regress A1 on all other variables
summary(lm(A1 ~ ., data = select(bfi, 01:05, A1:A5)))
```

Call:

```
lm(formula = A1 ~ ., data = select(bfi, 01:05, A1:A5))
```

Residuals:

	Min	1Q	Median	3Q	Max
	-3.5113	-0.9180	-0.3207	0.7330	4.5763

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	3.95198	0.20993	18.825	< 2e-16 ***
01	0.12114	0.02488	4.869	1.19e-06 ***
02	0.07446	0.01761	4.228	2.44e-05 ***
03	0.04967	0.02429	2.045	0.04098 *
04	-0.06568	0.02182	-3.010	0.00263 **
05	0.08494	0.02114	4.018	6.03e-05 ***
A2	-0.32318	0.02565	-12.599	< 2e-16 ***
A3	-0.13979	0.02469	-5.662	1.66e-08 ***
A4	-0.02005	0.01885	-1.063	0.28765
A5	-0.03240	0.02404	-1.348	0.17780

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.295 on 2637 degrees of freedom

(153 observations deleted due to missingness)

Multiple R-squared: 0.1541, Adjusted R-squared: 0.1512

F-statistic: 53.37 on 9 and 2637 DF, p-value: < 2.2e-16

```
# E.g., Regress 01 on all other variables
summary(lm(01 ~ ., data = select(bfi, 01:05, A1:A5)))
```

Call:

```
lm(formula = O1 ~ ., data = select(bfi, O1:O5, A1:A5))
```

Residuals:

Min	1Q	Median	3Q	Max
-4.5570	-0.5727	0.1064	0.7152	2.6925

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	2.730613	0.165923	16.457	< 2e-16	***
O2	-0.080779	0.013678	-5.906	3.96e-09	***
O3	0.264794	0.018230	14.525	< 2e-16	***
O4	0.098933	0.016922	5.846	5.64e-09	***
O5	-0.082739	0.016445	-5.031	5.20e-07	***
A1	0.073559	0.015107	4.869	1.19e-06	***
A2	0.052696	0.020556	2.564	0.01042	*
A3	0.032197	0.019345	1.664	0.09616	.
A4	0.008584	0.014692	0.584	0.55911	
A5	0.049964	0.018712	2.670	0.00763	**

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.009 on 2637 degrees of freedom

(153 observations deleted due to missingness)

Multiple R-squared: 0.2031, Adjusted R-squared: 0.2004

F-statistic: 74.67 on 9 and 2637 DF, p-value: < 2.2e-16

A computational short cut is to use the inverse of the correlation matrix

```
1 - 1 / diag(solve(corr_oa))
```

	O1	O2	O3	O4	O5	A1	A2	A3
	0.2030977	0.1625871	0.2766737	0.0766234	0.1910893	0.1540951	0.3401059	0.4002928
	A4	A5						
	0.1885574	0.3138769						

Iterative method

Here is sample code for principal axis factoring (using `eigen`)

```

rcorr_oa <- corr_oa # initialize reduced correlation
diag(rcorr_oa) <- 1 - 1 / diag(solve(corr_oa)) # replace communality
q <- 2 # number of factors
eps <- .0001 # convergence criterion
while (TRUE) {
  eigen_r <- eigen(rcorr_oa, symmetric = TRUE)
  lambdat <- t(eigen_r$vectors[, seq_len(q)]) * sqrt(eigen_r$values[seq_len(q)])
  # communality = sum of squared loadings
  new_h <- colSums(lambdat^2)
  if (max(abs(new_h - diag(rcorr_oa))) > eps) {
    diag(rcorr_oa) <- new_h
  } else {
    break # stop when convergence is met
  }
}
# Estimates with the iterative method
diag(rcorr_oa)

```

	01	02	03	04	05	A1	A2
	0.28265008	0.23396855	0.43541770	0.08877585	0.29766819	0.13249392	0.45364581
A3		A4	A5				
	0.58815163	0.25098540	0.39616917				

The final estimates are different from the initial ones.

Loading Matrix

```

# From our for loop for PAF
t(lambdat)

```

	[,1]	[,2]
[1,]	0.3548181	-0.3959189
[2,]	-0.1755868	0.4506955
[3,]	0.4741739	-0.4589373
[4,]	0.1651205	-0.2480128
[5,]	-0.2778184	0.4695370
[6,]	-0.3486927	-0.1044201
[7,]	0.6303503	0.2372005
[8,]	0.7164531	0.2737558

```
[9,] 0.4173173 0.2771697
[10,] 0.6018286 0.1842694
```

```
# Compared to `psych::fa()`
fa_unrotated <- psych::fa(
  select(bfi, 01:05, A1:A5),
  nfactors = 2,
  rotate = "none",
  fm = "pa")
fa_unrotated$loadings
```

Loadings:

	PA1	PA2
01	0.356	-0.392
02	-0.166	0.442
03	0.472	-0.453
04	0.163	-0.244
05	-0.270	0.476
A1	-0.345	
A2	0.631	0.231
A3	0.705	0.271
A4	0.415	0.270
A5	0.601	0.187

	PA1	PA2
SS loadings	2.025	1.084
Proportion Var	0.202	0.108
Cumulative Var	0.202	0.311

The values are not exactly the same as there are differences in implementation algorithms, but they are pretty close.

Number of Factors

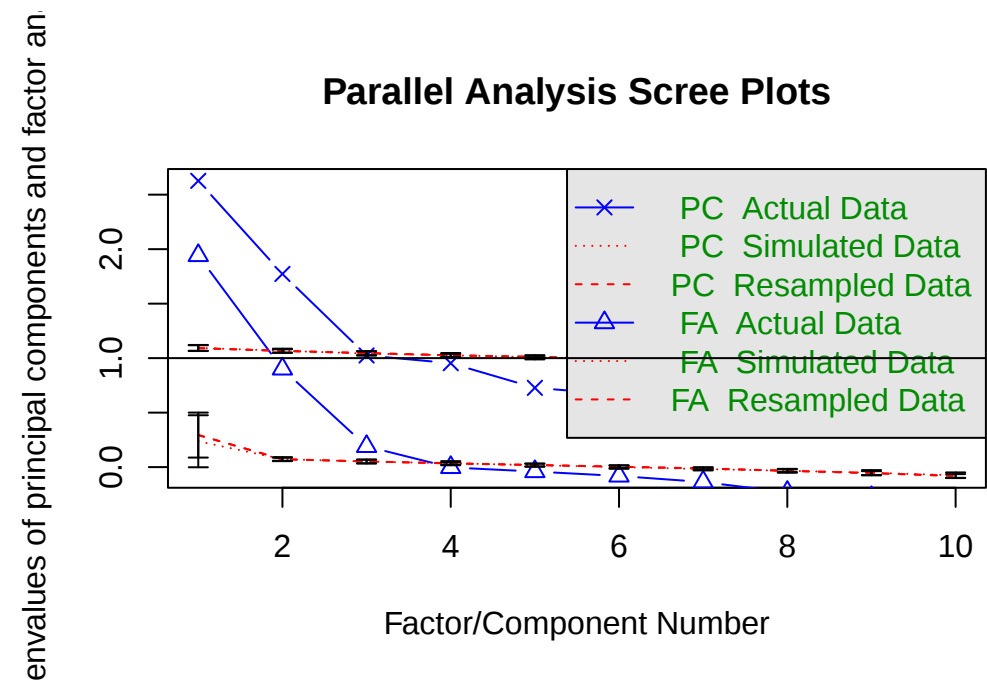
Parallel Analysis

```
fa.parallel(
  select(bfi, 01:05, A1:A5),
```

```

fm = "pa",
error.bars = TRUE
)

```



Parallel analysis suggests that the number of factors = 3 and the number of components = 2

It suggests 3 factors (with reduced correlation) and 2 components (with full correlation). In my experience, the number of components is usually more in line with theoretical expectations, but choosing either one would seem reasonable, and this should depend on theoretical considerations and interpretability of the solutions.

With polychoric correlation matrix

We need to first find the number of observations:

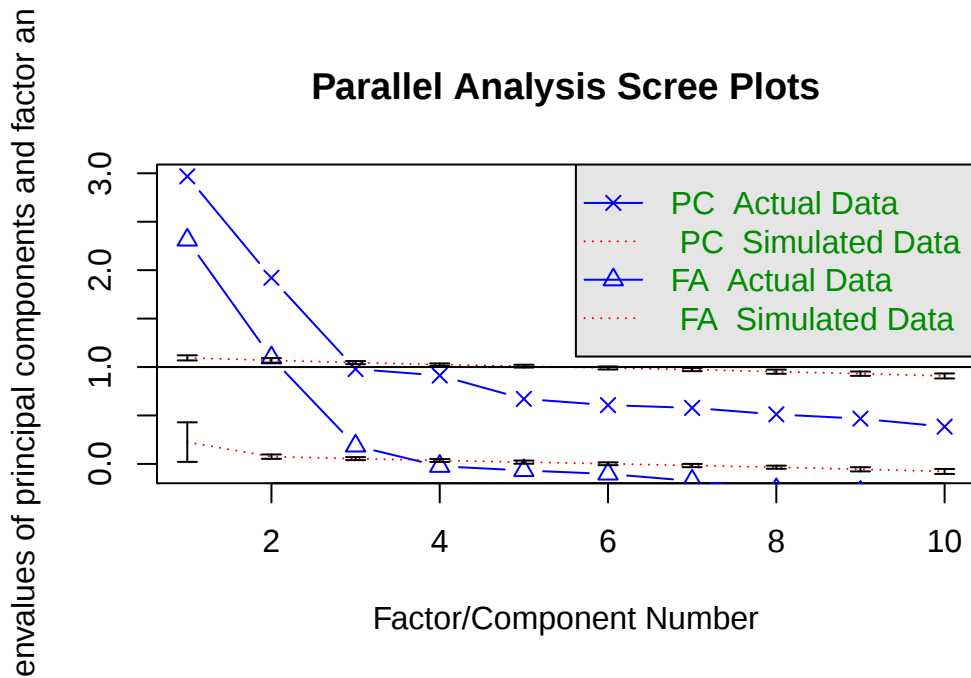
```

n_pcorr <-
  select(bfi, 01:05, A1:A5) |>
  drop_na() |>
  nrow()
n_pcorr

```

[1] 2647

```
psych::fa.parallel(  
  pcorr_oa,  
  n.obs = n_pcorr,  
  fm = "pa",  
  error.bars = TRUE  
)
```



Parallel analysis suggests that the number of factors = 3 and the number of components = 2

Hull method

```
# Reduced correlation  
HULL(select(bfi, 01:05, A1:A5), eigen_type = "SMC")
```

i Only CAF can be used as gof if method "PAF" is used. Setting gof to "CAF"

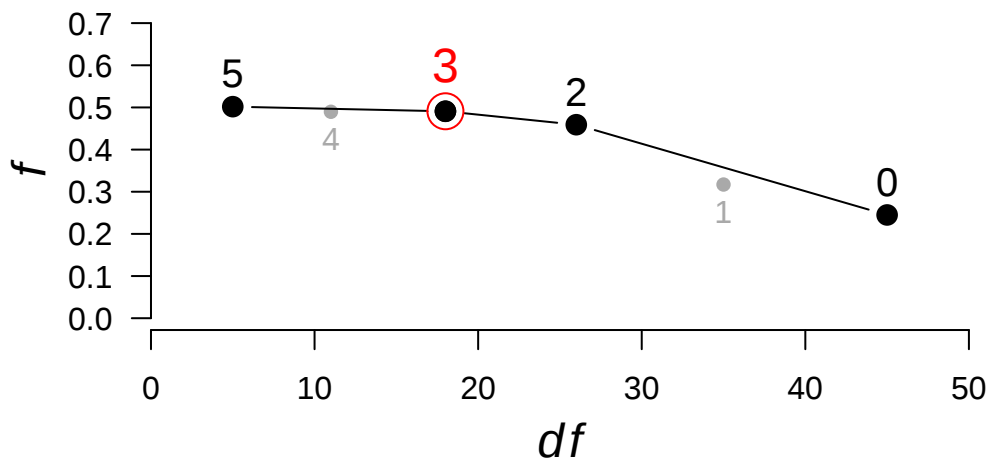
i 'x' was not a correlation matrix. Correlations are found from entered raw data.

Hull Analysis performed testing 0 to 5 factors.
PAF estimation and the CAF fit index was used.

-- Number of factors suggested by the Hull method -----

() With CAF: 3

Hull Method with PAF estimation and CAF



```
# Full correlation  
HULL(select(bfi, 01:05, A1:A5), eigen_type = "PCA")
```

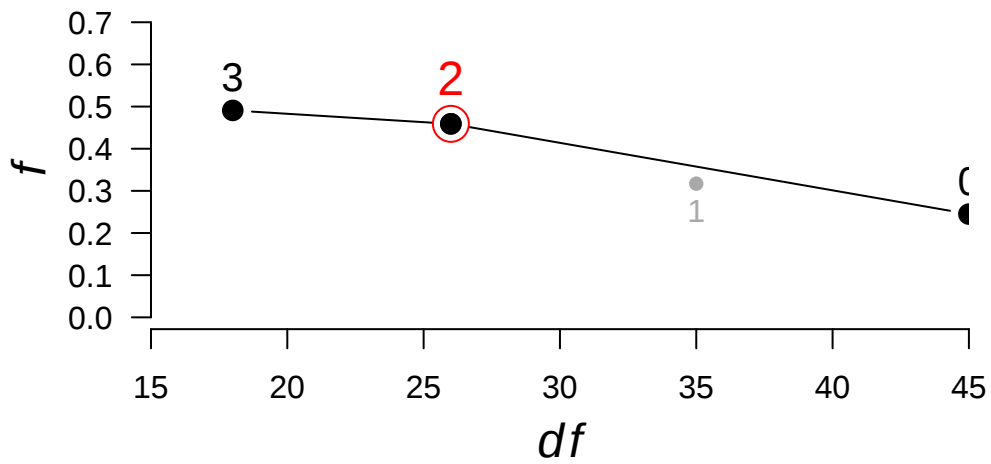
i Only CAF can be used as gof if method "PAF" is used. Setting gof to "CAF"
i 'x' was not a correlation matrix. Correlations are found from entered raw data.

Hull Analysis performed testing 0 to 3 factors.
PAF estimation and the CAF fit index was used.

-- Number of factors suggested by the Hull method -----

() With CAF: 2

Hull Method with PAF estimation and CAF



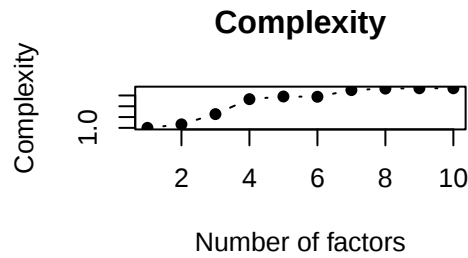
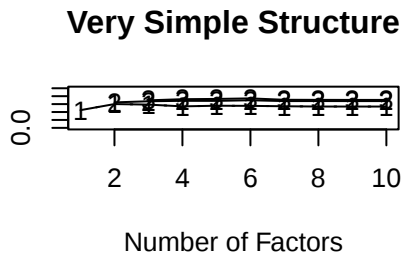
The conclusion is similar to parallel analysis.

Other methods

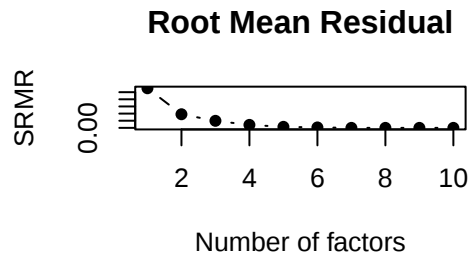
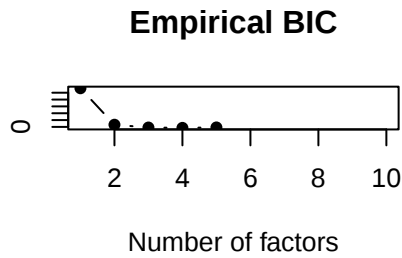
These should be examined as well. Note in this example, using `fm = "pa"` results in an error, so the suggestions are not based on PAF extraction.

```
psych::nfactors(select(bfi, 01:05, A1:A5))
```

Very Simple Structure Fit



Empirical BIC



Number of factors

Call: `vss(x = x, n = n, rotate = rotate, diagonal = diagonal, fm = fm, n.obs = n.obs, plot = FALSE, title = title, use = use, cor = cor)`

VSS complexity 1 achieves a maximum of 0.6 with 2 factors

VSS complexity 2 achieves a maximum of 0.7 with 6 factors

The Velicer MAP achieves a minimum of 0.03 with 2 factors

Empirical BIC achieves a minimum of -69.58 with 4 factors

Sample Size adjusted BIC achieves a minimum of -27.76 with 4 factors

Statistics by number of factors

	vss1	vss2	map	dof	chisq	prob	sqresid	fit	RMSEA	BIC	SABIC	complex
1	0.45	0.00	0.036	35	1.6e+03	1.3e-311	7.9	0.45	0.126	1310	1421.4	1.0
2	0.60	0.64	0.032	26	3.2e+02	5.4e-53	5.1	0.64	0.064	117	199.1	1.1
3	0.59	0.67	0.052	18	9.5e+01	1.7e-12	4.3	0.69	0.039	-48	9.4	1.3
4	0.54	0.68	0.087	11	2.5e+01	1.0e-02	3.9	0.73	0.021	-63	-27.8	1.5
5	0.56	0.68	0.134	5	2.8e+00	7.2e-01	3.7	0.74	0.000	-37	-21.0	1.6
6	0.56	0.70	0.213	0	5.5e-01	NA	3.4	0.76	NA	NA	NA	1.6
7	0.55	0.68	0.314	-4	1.6e-07	NA	3.5	0.76	NA	NA	NA	1.7
8	0.54	0.68	0.650	-7	2.2e-07	NA	3.5	0.75	NA	NA	NA	1.7
9	0.54	0.68	1.000	-9	0.0e+00	NA	3.5	0.75	NA	NA	NA	1.7
10	0.54	0.68	NA	-10	0.0e+00	NA	3.5	0.75	NA	NA	NA	1.7
	eChisq		SRMR		eCRMS	eBIC						

1	3.1e+03	1.1e-01	0.1259	2827
2	3.7e+02	3.8e-02	0.0502	160
3	9.8e+01	2.0e-02	0.0312	-45
4	1.8e+01	8.4e-03	0.0170	-70
5	2.1e+00	2.9e-03	0.0086	-38
6	3.7e-01	1.2e-03	NA	NA
7	1.2e-07	6.8e-07	NA	NA
8	1.4e-07	7.5e-07	NA	NA
9	1.1e-14	2.1e-10	NA	NA
10	1.1e-14	2.1e-10	NA	NA

Based on theoretical consideration, two factors seem reasonable.

Rotation

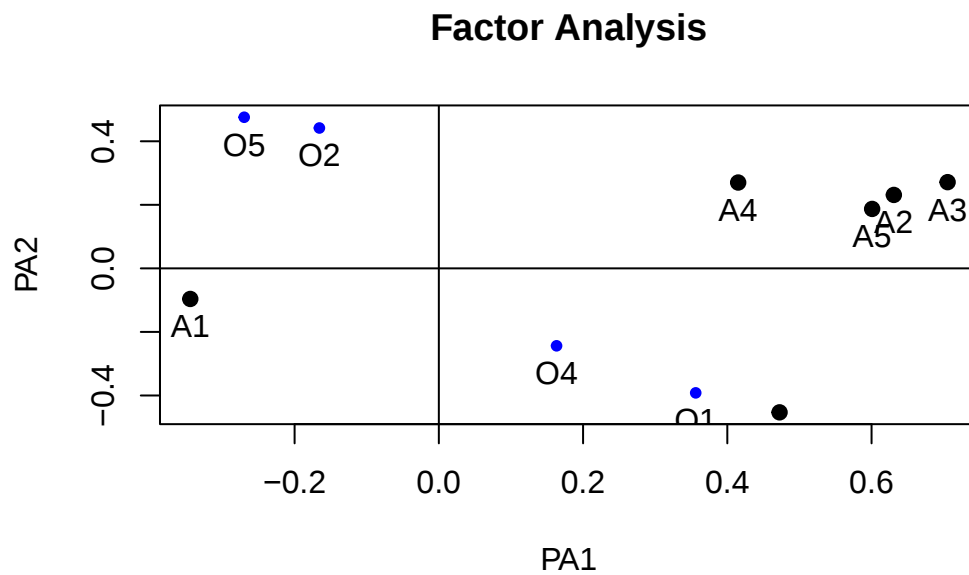
Complexity

Caution

The computation of variable complexity and factor complexity here are just for demonstrative purpose. You should not choose a rotation method based on these, because different rotations have different goals and complexity functions.

Unrotated solution

```
plot(fa_unrotated, labels = c(paste0("O", 1:5), paste0("A", 1:5)))
```



```
# Variable Complexity
lambda_unrotated <- fa_unrotated$loadings
sum(lambda_unrotated[, 1]^2 * lambda_unrotated[, 2]^2)
```

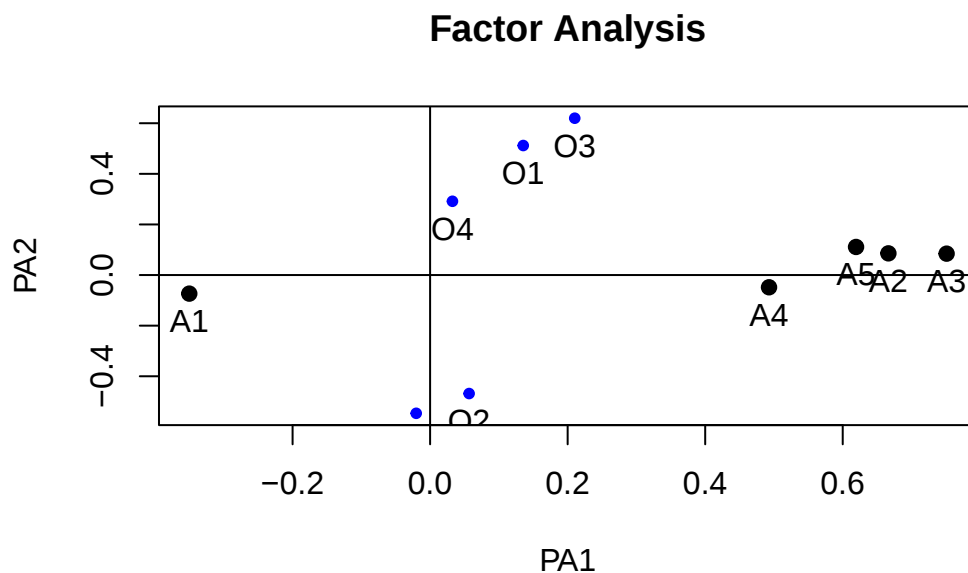
```
[1] 0.172905
```

```
# Factor Complexity
fac_complexity <- 0
for (j in seq_len(ncol(lambda_unrotated))) {
  for (i in seq_len(nrow(lambda_unrotated))) {
    fac_complexity <- fac_complexity + sum(lambda_unrotated[i, j]^2 * lambda_unrotated[i, j]^2)
  }
}
fac_complexity
```

```
[1] 4.447726
```

Varimax

```
fa_varimax <- psych::fa(
  select(bfi, O1:O5, A1:A5),
  nfactors = 2, fm = "pa", rotate = "varimax")
plot(fa_varimax, labels = c(paste0("O", 1:5), paste0("A", 1:5)))
```



```
# Variable Complexity
(lambda_varimax <- fa_varimax$loadings)
```

Loadings:

	PA1	PA2
O1	0.135	0.512
O2		-0.468
O3	0.210	0.620
O4		0.292
O5		-0.546
A1	-0.350	
A2	0.666	
A3	0.751	
A4	0.493	
A5	0.619	0.111

	PA1	PA2
SS loadings	1.825	1.284
Proportion Var	0.182	0.128
Cumulative Var	0.182	0.311

```
sum(lambda_varimax[, 1]^2 * lambda_varimax[, 2]^2)
```

```
[1] 0.03597018
```

```
# Factor Complexity
fac_complexity <- 0
for (j in seq_len(ncol(lambda_varimax))) {
  for (i in seq_len(nrow(lambda_varimax))) {
    fac_complexity <- fac_complexity + sum(lambda_varimax[i, j]^2 * lambda_varimax[-i, j]^2)
  }
}
fac_complexity
```

```
[1] 3.877532
```

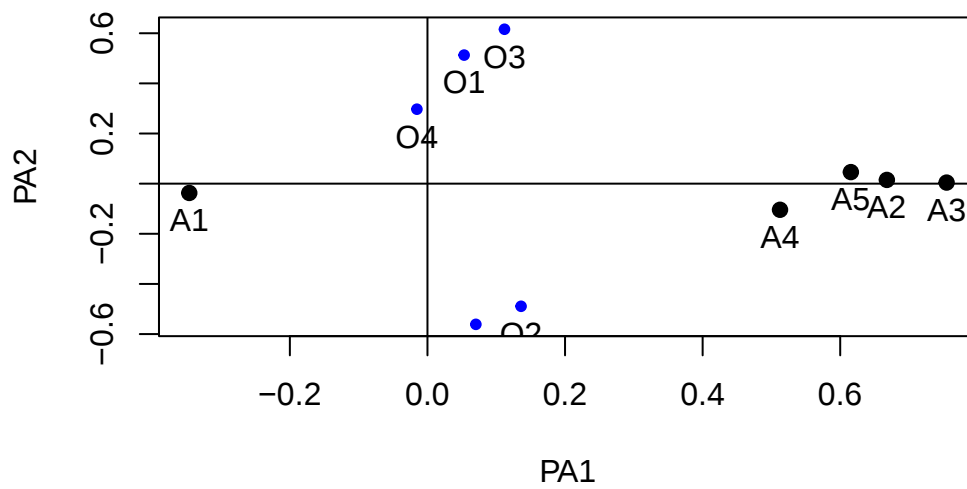
Oblimin

```
fa_oblimin <- psych::fa(
  select(bfi, 01:05, A1:A5),
  nfactors = 2, fm = "pa", rotate = "oblimin")
```

Loading required namespace: GPArotation

```
plot(fa_oblimin, labels = c(paste0("O", 1:5), paste0("A", 1:5)))
```

Factor Analysis



```
# Variable Complexity
(lambda_oblimin <- fa_oblimin$loadings)
```

Loadings:

	PA1	PA2
O1		0.513
O2	0.136	-0.489
O3	0.112	0.616
O4		0.297
O5		-0.561
A1	-0.346	
A2	0.668	
A3	0.755	
A4	0.512	-0.104
A5	0.615	

	PA1	PA2
SS loadings	1.816	1.300
Proportion Var	0.182	0.130
Cumulative Var	0.182	0.312

```
sum(lambda_oblimin[, 1]^2 * lambda_oblimin[, 2]^2)
```

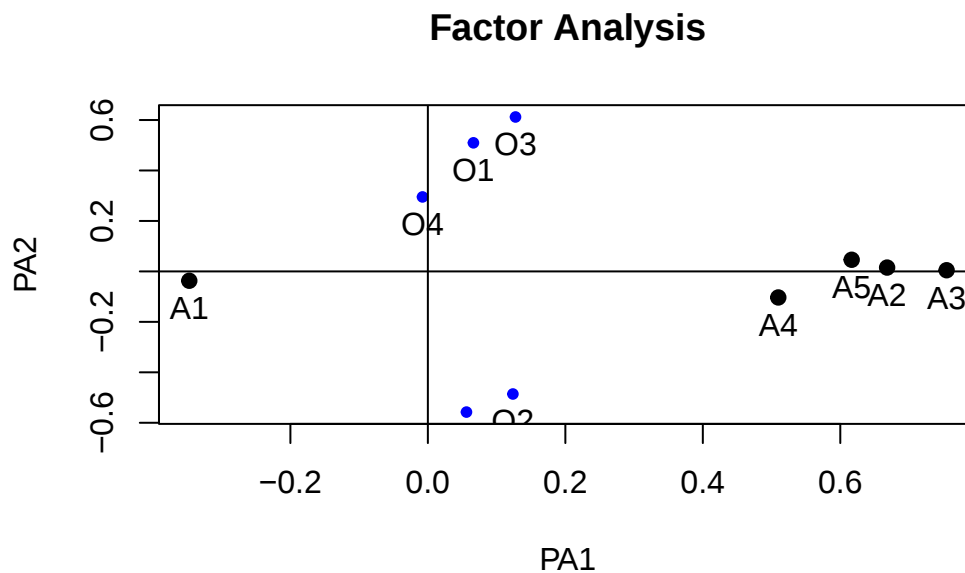
```
[1] 0.01545015
```

```
# Factor Complexity
fac_complexity <- 0
for (j in seq_len(ncol(lambda_oblimin))) {
  for (i in seq_len(nrow(lambda_oblimin))) {
    fac_complexity <- fac_complexity + sum(lambda_oblimin[i, j]^2 * lambda_oblimin[-i, j]^2)
  }
}
fac_complexity
```

```
[1] 3.858235
```

Geomin (Oblique)

```
fa_geomin <- psych::fa(
  select(bfi, 01:05, A1:A5),
  nfactors = 2, fm = "pa", rotate = "geominQ")
plot(fa_geomin, labels = c(paste0("O", 1:5), paste0("A", 1:5)))
```



```
# Variable Complexity
(lambda_geomin <- fa_geomin$loadings)
```

Loadings:

	PA1	PA2
O1		0.510
O2	0.124	-0.486
O3	0.128	0.612
O4		0.295
O5		-0.558
A1	-0.347	
A2	0.668	
A3	0.755	
A4	0.510	-0.103
A5	0.617	

	PA1	PA2
SS loadings	1.816	1.283
Proportion Var	0.182	0.128
Cumulative Var	0.182	0.310

```
sum(lambda_geomin[, 1]^2 * lambda_geomin[, 2]^2)
```

```
[1] 0.01569736
```

```
# Factor Complexity
fac_complexity <- 0
for (j in seq_len(ncol(lambda_geomin))) {
  for (i in seq_len(nrow(lambda_geomin))) {
    fac_complexity <- fac_complexity + sum(lambda_geomin[i, j]^2 * lambda_geomin[-i, j]^2)
  }
}
fac_complexity
```

```
[1] 3.823991
```

Target Rotation

In practice, we usually have some idea about the factor structure, and we can rotate the loadings to something close to that idea. This is an example of blurring the distinction between EFA and CFA.

Let's say we think that the five O items are for one factor, and the five A items are for another.

```
# We want the first five items to load on factor 1 but not on factor 2,
# and the reverse for the next five items
target_mat <- matrix(
  c(NA, NA, NA, NA, NA, 0, 0, 0, 0, 0,
    0, 0, 0, 0, 0, NA, NA, NA, NA, NA),
  nrow = 10
)
fa_target <- psych::fa(
  select(bfi, 01:05, A1:A5),
  nfactors = 2, fm = "pa",
  rotate = "targetQ", Target = target_mat)
fa_target
```

Factor Analysis using method = pa

Call: psych::fa(r = select(bfi, 01:05, A1:A5), nfactors = 2, rotate = "targetQ",
fm = "pa", Target = target_mat)

Standardized loadings (pattern matrix) based upon correlation matrix

	PA2	PA1	h2	u2	com
01	0.06	0.51	0.280	0.72	1.0
02	0.13	-0.49	0.223	0.78	1.1
03	0.12	0.61	0.428	0.57	1.1
04	-0.01	0.30	0.086	0.91	1.0
05	0.06	-0.56	0.299	0.70	1.0
A1	-0.35	-0.04	0.128	0.87	1.0
A2	0.67	0.02	0.451	0.55	1.0
A3	0.75	0.01	0.571	0.43	1.0
A4	0.51	-0.10	0.245	0.75	1.1
A5	0.62	0.05	0.396	0.60	1.0

	PA2	PA1
SS loadings	1.81	1.29
Proportion Var	0.18	0.13
Cumulative Var	0.18	0.31
Proportion Explained	0.58	0.42

Cumulative Proportion 0.58 1.00

With factor correlations of
PA2 PA1
PA2 1.00 0.25
PA1 0.25 1.00

Mean item complexity = 1
Test of the hypothesis that 2 factors are sufficient.

df null model = 45 with the objective function = 1.58 with Chi Square = 4409.68
df of the model are 26 and the objective function was 0.12

The root mean square of the residuals (RMSR) is 0.04
The df corrected root mean square of the residuals is 0.05

The harmonic n.obs is 2766 with the empirical chi square 362.29 with prob < 6e-61
The total n.obs was 2800 with Likelihood Chi Square = 323 with prob < 5.2e-53

Tucker Lewis Index of factoring reliability = 0.882
RMSEA index = 0.064 and the 90 % confidence intervals are 0.058 0.07
BIC = 116.63
Fit based upon off diagonal values = 0.97
Measures of factor score adequacy

	PA2	PA1
Correlation of (regression) scores with factors	0.88	0.81
Multiple R square of scores with factors	0.77	0.66
Minimum correlation of possible factor scores	0.54	0.32

Percentage of Variance Accounted For

Let's use oblimin solution

```
# Proportion of variance by the first two eigenvalues
ev_oa <- fa_oblimin$e.values
sum(ev_oa[1:2]) / sum(ev_oa)
```

```
[1] 0.4399516
```

The first two eigenvalues account for 44% of the total variance in the data.

Correlation Residual

This does not depend on rotation method.

```
round(fa_unrotated$residual, digits = 2)
```

	01	02	03	04	05	A1	A2	A3	A4	A5
01	0.72	0.02	0.05	0.02	0.04	0.10	0.00	0.00	0.02	0.02
02	0.02	0.78	0.02	0.07	0.07	0.06	0.02	0.00	-0.01	0.02
03	0.05	0.02	0.57	0.01	0.03	0.06	-0.03	0.01	0.00	0.04
04	0.02	0.07	0.01	0.91	-0.02	-0.04	0.04	-0.01	-0.04	-0.04
05	0.04	0.07	0.03	-0.02	0.70	0.07	-0.03	0.01	0.00	0.02
A1	0.10	0.06	0.06	-0.04	0.07	0.87	-0.10	0.00	0.02	0.04
A2	0.00	0.02	-0.03	0.04	-0.03	-0.10	0.55	-0.02	0.01	-0.03
A3	0.00	0.00	0.01	-0.01	0.01	0.00	-0.02	0.43	-0.01	0.03
A4	0.02	-0.01	0.00	-0.04	0.00	0.02	0.01	-0.01	0.75	0.01
A5	0.02	0.02	0.04	-0.04	0.02	0.04	-0.03	0.03	0.01	0.60

As a general rule of thumb, researchers should pay attention to residuals with absolute value $> .10$. In our case, there were a few values close to $.10$, but it doesn't seem to be strong evidence that there may be an additional factor needed.

Confidence Intervals

The loadings in EFA are quite unstable, especially with rotation. We can obtain confidence intervals using bootstrap, which is available in `psych::fa()` by specifying

```
fa_oblimin_boot <- psych::fa(  
  select(bfi, 01:05, A1:A5),  
  nfactors = 2,  
  n.iter = 200, # number of bootstrap samples  
  fm = "pa", rotate = "oblimin")  
fa_oblimin_boot
```

```
Factor Analysis with confidence intervals using method = psych::fa(r = select(bfi, 01:05, A1  
  rotate = "oblimin", fm = "pa")  
Factor Analysis using method = pa  
Call: psych::fa(r = select(bfi, 01:05, A1:A5), nfactors = 2, n.iter = 200,  
  rotate = "oblimin", fm = "pa")
```

Standardized loadings (pattern matrix) based upon correlation matrix

	PA1	PA2	h2	u2	com
O1	0.05	0.51	0.280	0.72	1.0
O2	0.14	-0.49	0.223	0.78	1.2
O3	0.11	0.62	0.428	0.57	1.1
O4	-0.02	0.30	0.086	0.91	1.0
O5	0.07	-0.56	0.299	0.70	1.0
A1	-0.35	-0.04	0.128	0.87	1.0
A2	0.67	0.01	0.451	0.55	1.0
A3	0.75	0.00	0.571	0.43	1.0
A4	0.51	-0.10	0.245	0.75	1.1
A5	0.62	0.05	0.396	0.60	1.0

	PA1	PA2
SS loadings	1.81	1.30
Proportion Var	0.18	0.13
Cumulative Var	0.18	0.31
Proportion Explained	0.58	0.42
Cumulative Proportion	0.58	1.00

With factor correlations of

	PA1	PA2
PA1	1.00	0.26
PA2	0.26	1.00

Mean item complexity = 1

Test of the hypothesis that 2 factors are sufficient.

df null model = 45 with the objective function = 1.58 with Chi Square = 4409.68
df of the model are 26 and the objective function was 0.12

The root mean square of the residuals (RMSR) is 0.04

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The harmonic n.obs is 2766 with the empirical chi square 362.29 with prob < 6e-61

The total n.obs was 2800 with Likelihood Chi Square = 323 with prob < 5.2e-53

Tucker Lewis Index of factoring reliability = 0.882

RMSEA index = 0.064 and the 90 % confidence intervals are 0.058 0.07

BIC = 116.63

Fit based upon off diagonal values = 0.97

Measures of factor score adequacy

PA1	PA2
-----	-----

Correlation of (regression) scores with factors	0.88	0.81
Multiple R square of scores with factors	0.77	0.66
Minimum correlation of possible factor scores	0.54	0.33

Coefficients and bootstrapped confidence intervals

	low	PA1	upper	low	PA2	upper
O1	0.02	0.05	0.10	0.46	0.51	0.56
O2	0.10	0.14	0.17	-0.54	-0.49	-0.44
O3	0.08	0.11	0.15	0.56	0.62	0.67
O4	-0.06	-0.02	0.03	0.24	0.30	0.35
O5	0.03	0.07	0.11	-0.61	-0.56	-0.51
A1	-0.39	-0.35	-0.30	-0.10	-0.04	0.01
A2	0.63	0.67	0.71	-0.02	0.01	0.05
A3	0.72	0.75	0.78	-0.02	0.00	0.04
A4	0.47	0.51	0.55	-0.15	-0.10	-0.07
A5	0.57	0.62	0.66	0.00	0.05	0.09

Interfactor correlations and bootstrapped confidence intervals

	lower	estimate	upper
PA1-PA2	0.21	0.26	0.32

In practice, you should use more bootstrap samples (e.g., 1999). We can do the same for target rotation:

```
fa_target_boot <- psych::fa(
  select(bfi, O1:O5, A1:A5),
  nfactors = 2,
  n.iter = 200, # number of bootstrap samples
  fm = "pa",
  rotate = "targetQ", Target = target_mat)
fa_target_boot
```

```
Factor Analysis with confidence intervals using method = psych::fa(r = select(bfi, O1:O5, A1:A5),
  rotate = "targetQ", fm = "pa", Target = target_mat)
```

```
Factor Analysis using method = pa
```

```
Call: psych::fa(r = select(bfi, O1:O5, A1:A5), nfactors = 2, n.iter = 200,
  rotate = "targetQ", fm = "pa", Target = target_mat)
```

```
Standardized loadings (pattern matrix) based upon correlation matrix
```

	PA2	PA1	h2	u2	com
O1	0.06	0.51	0.280	0.72	1.0
O2	0.13	-0.49	0.223	0.78	1.1
O3	0.12	0.61	0.428	0.57	1.1

O4	-0.01	0.30	0.086	0.91	1.0
O5	0.06	-0.56	0.299	0.70	1.0
A1	-0.35	-0.04	0.128	0.87	1.0
A2	0.67	0.02	0.451	0.55	1.0
A3	0.75	0.01	0.571	0.43	1.0
A4	0.51	-0.10	0.245	0.75	1.1
A5	0.62	0.05	0.396	0.60	1.0

	PA2	PA1
SS loadings	1.81	1.29
Proportion Var	0.18	0.13
Cumulative Var	0.18	0.31
Proportion Explained	0.58	0.42
Cumulative Proportion	0.58	1.00

With factor correlations of

	PA2	PA1
PA2	1.00	0.25
PA1	0.25	1.00

Mean item complexity = 1
 Test of the hypothesis that 2 factors are sufficient.

df null model = 45 with the objective function = 1.58 with Chi Square = 4409.68
 df of the model are 26 and the objective function was 0.12

The root mean square of the residuals (RMSR) is 0.04
 The df corrected root mean square of the residuals is 0.05

The harmonic n.obs is 2766 with the empirical chi square 362.29 with prob < 6e-61
 The total n.obs was 2800 with Likelihood Chi Square = 323 with prob < 5.2e-53

Tucker Lewis Index of factoring reliability = 0.882
 RMSEA index = 0.064 and the 90 % confidence intervals are 0.058 0.07
 BIC = 116.63
 Fit based upon off diagonal values = 0.97
 Measures of factor score adequacy

	PA2	PA1
Correlation of (regression) scores with factors	0.88	0.81
Multiple R square of scores with factors	0.77	0.66
Minimum correlation of possible factor scores	0.54	0.32

Coefficients and bootstrapped confidence intervals

	low	PA2	upper	low	PA1	upper
O1	0.02	0.06	0.10	0.47	0.51	0.56
O2	0.09	0.13	0.16	-0.53	-0.49	-0.43
O3	0.08	0.12	0.16	0.57	0.61	0.66
O4	-0.06	-0.01	0.03	0.24	0.30	0.35
O5	0.02	0.06	0.10	-0.61	-0.56	-0.51
A1	-0.40	-0.35	-0.30	-0.08	-0.04	0.01
A2	0.63	0.67	0.71	-0.02	0.02	0.05
A3	0.72	0.75	0.79	-0.02	0.01	0.04
A4	0.47	0.51	0.55	-0.14	-0.10	-0.07
A5	0.58	0.62	0.66	0.01	0.05	0.09

Interfactor correlations and bootstrapped confidence intervals

	lower	estimate	upper
PA2-PA1	0.2	0.25	0.3

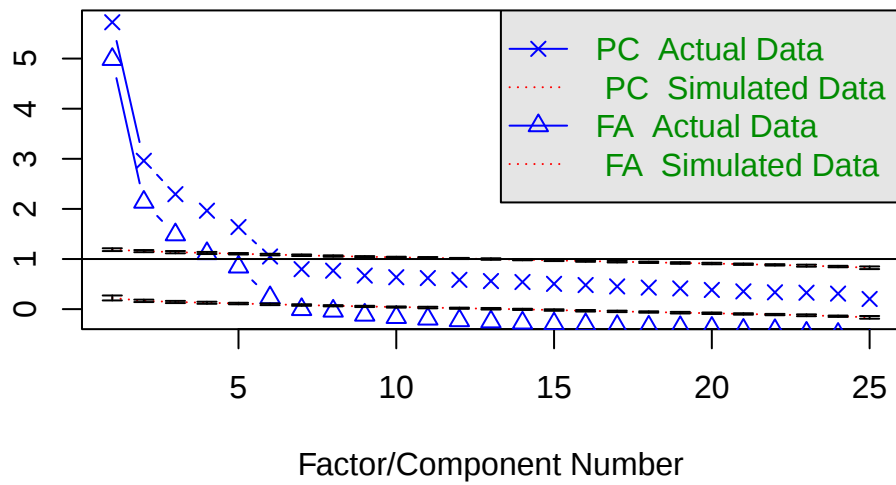
Full BFI

Target rotation with BFI, with polychoric correlation

```
pcorr_bfi <- lavaan::lavCor(select(bfi, A1:O5), ordered = TRUE)
n_pcorr_bfi <-
  select(bfi, A1:O5) |>
  drop_na() |>
  nrow()
# Parallel analysis
psych::fa.parallel(
  pcorr_bfi,
  n.obs = n_pcorr_bfi,
  fm = "pa",
  error.bars = TRUE
)
```

envalues of principal components and factor an

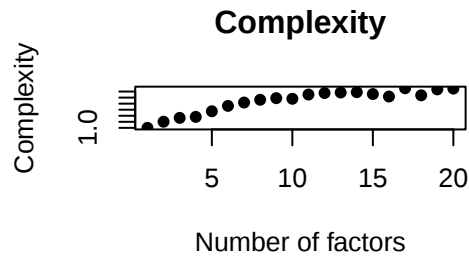
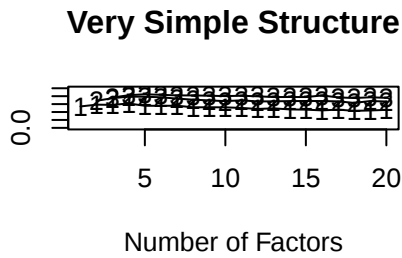
Parallel Analysis Scree Plots



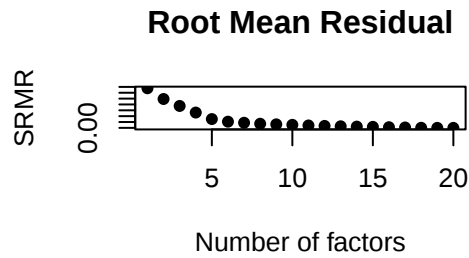
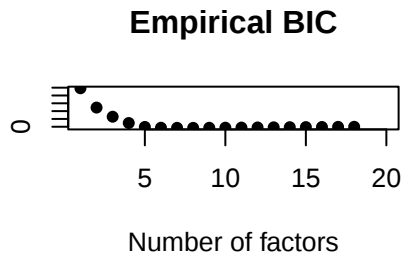
Parallel analysis suggests that the number of factors = 6 and the number of components = 5

```
# Other methods
n factors(pcorr_bfi, n.obs = n_pcorr_bfi)
```

Very Simple Structure Fit



Empirical BIC



Number of factors

Call: `vss(x = x, n = n, rotate = rotate, diagonal = diagonal, fm = fm, n.obs = n.obs, plot = FALSE, title = title, use = use, cor = cor)`

VSS complexity 1 achieves a maximum of Although the `vss.max` shows 4 factors, it is probal

VSS complexity 2 achieves a maximum of 0.79 with 4 factors

The Velicer MAP achieves a minimum of 0.02 with 5 factors

Empirical BIC achieves a minimum of -711.98 with 8 factors

Sample Size adjusted BIC achieves a minimum of -125.98 with 12 factors

Statistics by number of factors

	vss1	vss2	map	dof	chisq	prob	sqresid	fit	RMSEA	BIC	SABIC
1	0.54	0.00	0.032	275	1.4e+04	0.0e+00	27.3	0.54	0.1438	11980	12853.7
2	0.59	0.68	0.026	251	9.3e+03	0.0e+00	18.8	0.68	0.1213	7295	8092.2
3	0.58	0.75	0.023	228	6.6e+03	0.0e+00	13.9	0.77	0.1074	4860	5584.3
4	0.62	0.79	0.019	206	4.4e+03	0.0e+00	10.1	0.83	0.0912	2774	3428.8
5	0.56	0.79	0.016	185	2.2e+03	0.0e+00	7.6	0.87	0.0673	783	1370.9
6	0.56	0.77	0.017	165	1.4e+03	1.7e-188	6.6	0.89	0.0547	82	606.2
7	0.56	0.74	0.021	146	9.6e+02	1.8e-120	6.2	0.90	0.0479	-175	288.7
8	0.52	0.71	0.023	128	7.1e+02	1.2e-80	5.7	0.90	0.0430	-293	114.0
9	0.52	0.69	0.028	111	5.2e+02	5.4e-54	5.5	0.91	0.0388	-347	6.1
10	0.52	0.68	0.033	95	3.9e+02	1.0e-37	5.0	0.92	0.0358	-349	-47.0
11	0.50	0.70	0.040	80	2.5e+02	4.7e-20	5.0	0.92	0.0299	-369	-115.2

12	0.48	0.67	0.048	66	1.8e+02	2.3e-12	4.7	0.92	0.0265	-336	-126.0
13	0.49	0.67	0.058	53	1.3e+02	3.2e-08	4.5	0.92	0.0242	-285	-116.3
14	0.47	0.68	0.066	41	9.0e+01	1.8e-05	4.2	0.93	0.0220	-230	-99.9
15	0.46	0.68	0.081	30	6.0e+01	9.3e-04	3.7	0.94	0.0202	-174	-78.7
16	0.44	0.68	0.097	20	2.4e+01	2.5e-01	3.3	0.94	0.0089	-132	-68.5
17	0.46	0.67	0.114	11	7.8e+00	7.3e-01	3.7	0.94	0.0000	-78	-43.0
18	0.43	0.63	0.134	3	4.5e-01	9.3e-01	3.5	0.94	0.0000	-23	-13.4
19	0.44	0.68	0.157	-4	3.8e-04	NA	3.7	0.94	NA	NA	NA
20	0.45	0.67	0.199	-10	8.4e-05	NA	3.7	0.94	NA	NA	NA
complex eChisq SRMR eCRMS eBIC											
1	1.0	2.7e+04	1.4e-01	0.1414	24627						
2	1.2	1.4e+04	9.9e-02	0.1078	12266						
3	1.3	8.2e+03	7.5e-02	0.0859	6414						
4	1.4	4.0e+03	5.2e-02	0.0631	2392						
5	1.5	1.3e+03	3.0e-02	0.0383	-121						
6	1.7	6.6e+02	2.1e-02	0.0286	-631						
7	1.8	4.4e+02	1.7e-02	0.0249	-697						
8	1.9	2.9e+02	1.4e-02	0.0214	-712						
9	2.0	2.1e+02	1.2e-02	0.0198	-653						
10	2.0	1.6e+02	1.0e-02	0.0184	-585						
11	2.1	9.7e+01	8.2e-03	0.0158	-526						
12	2.2	6.4e+01	6.6e-03	0.0141	-451						
13	2.2	4.3e+01	5.4e-03	0.0129	-370						
14	2.2	2.7e+01	4.3e-03	0.0117	-292						
15	2.1	1.7e+01	3.5e-03	0.0109	-216						
16	2.0	7.6e+00	2.3e-03	0.0088	-148						
17	2.3	2.6e+00	1.3e-03	0.0070	-83						
18	2.1	1.0e-01	2.6e-04	0.0026	-23						
19	2.3	7.5e-05	7.2e-06	NA	NA						
20	2.3	1.8e-05	3.5e-06	NA	NA						

```
# Target rotation
target_mat_bfi <- matrix(0, nrow = 25, ncol = 5)
target_mat_bfi[1:5, 1] <- NA
target_mat_bfi[6:10, 2] <- NA
target_mat_bfi[11:15, 3] <- NA
target_mat_bfi[16:20, 4] <- NA
target_mat_bfi[21:25, 5] <- NA
fa_target_bfi <- psych::fa(
  pcorr_bfi, n.obs = n_pcorr_bfi,
  nfactors = 5, fm = "pa",
  rotate = "targetQ", Target = target_mat_bfi)
```

fa_target_bfi

Factor Analysis using method = pa

Call: psych::fa(r = pcorr_bfi, nfactors = 5, n.obs = n_pcorr_bfi, rotate = "targetQ",
fm = "pa", Target = target_mat_bfi)

Standardized loadings (pattern matrix) based upon correlation matrix

	PA4	PA3	PA2	PA1	PA5	h2	u2	com
A1	0.15	0.18	0.07	-0.51	-0.09	0.26	0.74	1.5
A2	0.03	0.05	0.07	0.69	0.03	0.54	0.46	1.0
A3	0.02	0.18	0.02	0.69	0.04	0.61	0.39	1.1
A4	-0.02	0.10	0.21	0.47	-0.18	0.35	0.65	1.9
A5	-0.11	0.28	0.00	0.55	0.04	0.53	0.47	1.6
C1	0.07	-0.02	0.59	0.00	0.17	0.40	0.60	1.2
C2	0.18	-0.05	0.71	0.06	0.06	0.51	0.49	1.2
C3	0.04	-0.08	0.61	0.10	-0.07	0.36	0.64	1.1
C4	0.17	0.03	-0.69	0.01	-0.03	0.55	0.45	1.1
C5	0.21	-0.10	-0.60	0.01	0.12	0.49	0.51	1.4
E1	-0.01	-0.62	0.13	-0.07	-0.02	0.39	0.61	1.1
E2	0.17	-0.71	-0.01	-0.04	0.01	0.61	0.39	1.1
E3	0.07	0.51	-0.02	0.22	0.27	0.50	0.50	2.0
E4	-0.05	0.64	0.00	0.27	-0.13	0.60	0.40	1.4
E5	0.13	0.50	0.28	0.01	0.17	0.46	0.54	2.0
N1	0.86	0.21	0.01	-0.19	-0.08	0.74	0.26	1.2
N2	0.81	0.14	0.02	-0.16	0.00	0.66	0.34	1.1
N3	0.77	0.00	-0.03	0.02	0.01	0.59	0.41	1.0
N4	0.56	-0.33	-0.13	0.08	0.13	0.55	0.45	1.9
N5	0.56	-0.17	0.00	0.20	-0.15	0.40	0.60	1.6
O1	0.00	0.18	0.06	0.00	0.55	0.39	0.61	1.3
O2	0.20	0.02	-0.08	0.15	-0.50	0.30	0.70	1.6
O3	0.02	0.27	-0.01	0.05	0.63	0.54	0.46	1.4
O4	0.19	-0.28	-0.04	0.19	0.48	0.35	0.65	2.4
O5	0.11	0.05	-0.04	0.04	-0.59	0.36	0.64	1.1

	PA4	PA3	PA2	PA1	PA5
SS loadings	2.96	2.61	2.33	2.26	1.85
Proportion Var	0.12	0.10	0.09	0.09	0.07
Cumulative Var	0.12	0.22	0.32	0.41	0.48
Proportion Explained	0.25	0.22	0.19	0.19	0.15
Cumulative Proportion	0.25	0.46	0.66	0.85	1.00

With factor correlations of

	PA4	PA3	PA2	PA1	PA5
--	-----	-----	-----	-----	-----

```

PA4  1.00 -0.21 -0.21 -0.09 0.02
PA3 -0.21  1.00  0.29  0.34 0.16
PA2 -0.21  0.29  1.00  0.23 0.20
PA1 -0.09  0.34  0.23  1.00 0.16
PA5  0.02  0.16  0.20  0.16 1.00

```

Mean item complexity = 1.4

Test of the hypothesis that 5 factors are sufficient.

df null model = 300 with the objective function = 9.59 with Chi Square = 23261.83
df of the model are 185 and the objective function was 0.92

The root mean square of the residuals (RMSR) is 0.03

The df corrected root mean square of the residuals is 0.04

The harmonic n.obs is 2436 with the empirical chi square 1322.13 with prob < 8.1e-171

The total n.obs was 2436 with Likelihood Chi Square = 2226.01 with prob < 0

Tucker Lewis Index of factoring reliability = 0.856

RMSEA index = 0.067 and the 90 % confidence intervals are 0.065 0.07

BIC = 783.36

Fit based upon off diagonal values = 0.98

Measures of factor score adequacy

	PA4	PA3	PA2	PA1	PA5
Correlation of (regression) scores with factors	0.94	0.92	0.90	0.90	0.87
Multiple R square of scores with factors	0.89	0.84	0.81	0.81	0.76
Minimum correlation of possible factor scores	0.77	0.68	0.63	0.62	0.52

Using the lavaan R package

More information at <https://lavaan.ugent.be/tutorial/efa.html>

```

# One- to Six-factor models
efa_fit <- efa(
  select(bfi, A1:O5),
  nfactors = 1:6,
  ordered = TRUE # only needed for ordinal data
)
summary(efa_fit)

```

This is lavaan 0.6-19 -- running exploratory factor analysis

Estimator	DWLS
Rotation method	GEOMIN OBLIQUE
Geomin epsilon	0.001
Rotation algorithm (rstarts)	GPA (30)
Standardized metric	TRUE
Row weights	None

	Used	Total
Number of observations	2436	2800

Overview models:

	chisq	df	pvalue	cfi	rmsea
nfactors = 1	16233.810	275	0	0.363	0.148
nfactors = 2	10653.913	251	0	0.609	0.122
nfactors = 3	7391.517	228	0	0.723	0.107
nfactors = 4	4711.331	206	0	0.819	0.091
nfactors = 5	1856.317	185	0	0.914	0.066
nfactors = 6	1058.461	165	0	0.950	0.054

Eigenvalues correlation matrix:

ev1	ev2	ev3	ev4	ev5	ev6	ev7	ev8	ev9	ev10
5.725	2.960	2.294	1.964	1.638	1.050	0.797	0.767	0.669	0.637
ev11	ev12	ev13	ev14	ev15	ev16	ev17	ev18	ev19	ev20
0.620	0.584	0.560	0.541	0.504	0.480	0.453	0.429	0.413	0.383
ev21	ev22	ev23	ev24	ev25					
0.354	0.335	0.328	0.310	0.203					

Number of factors: 1

Standardized loadings: (* = significant at 1% level)

	f1	unique.var	communalities
A1	.*	0.930	0.070
A2	-0.503*	0.747	0.253
A3	-0.587*	0.655	0.345
A4	-0.432*	0.813	0.187
A5	-0.623*	0.611	0.389
C1	-0.377*	0.858	0.142
C2	-0.360*	0.871	0.129
C3	-0.330*	0.891	0.109
C4	0.510*	0.740	0.260

C5	0.510*	0.740	0.260
E1	0.418*	0.825	0.175
E2	0.626*	0.609	0.391
E3	-0.542*	0.706	0.294
E4	-0.624*	0.610	0.390
E5	-0.497*	0.753	0.247
N1	0.701*	0.508	0.492
N2	0.674*	0.546	0.454
N3	0.547*	0.700	0.300
N4	0.559*	0.688	0.312
N5	0.407*	0.834	0.166
O1	-0.355*	0.874	0.126
O2	.*	0.947	0.053
O3	-0.412*	0.830	0.170
O4	*	0.997	0.003
O5	.*	0.947	0.053

	f1
Sum of squared loadings	5.769
Proportion of total	1.000
Proportion var	0.231
Cumulative var	0.231

Number of factors: 2

Standardized loadings: (* = significant at 1% level)

	f1	f2	unique.var	communalities
A1	.*	*	0.925	0.075
A2	0.591*		0.656	0.344
A3	0.672*		0.554	0.446
A4	0.441*	*	0.784	0.216
A5	0.641*	.*	0.558	0.442
C1	0.431*		0.811	0.189
C2	0.451*		0.800	0.200
C3	0.338*	*	0.869	0.131
C4	-0.420*	.*	0.707	0.293
C5	-0.376*	0.343*	0.700	0.300
E1	-0.467*		0.776	0.224
E2	-0.579*	.*	0.570	0.430
E3	0.663*	*	0.571	0.429
E4	0.624*	.*	0.563	0.437
E5	0.615*	*	0.629	0.371

N1		0.859*	0.264	0.736
N2		0.845*	0.292	0.708
N3		0.750*	0.437	0.563
N4	.*	0.614*	0.539	0.461
N5	*	0.525*	0.705	0.295
O1	0.456*	*	0.797	0.203
O2	.*	*	0.944	0.056
O3	0.542*	.*	0.713	0.287
O4	.*	0.304*	0.902	0.098
O5	.*		0.925	0.075

	f1	f2	total
Sum of sq (obliq) loadings	4.821	3.188	8.009
Proportion of total	0.602	0.398	1.000
Proportion var	0.193	0.128	0.320
Cumulative var	0.193	0.320	0.320

Factor correlations: (* = significant at 1% level)

	f1	f2
f1	1.000	
f2	-0.161*	1.000

Number of factors: 3

Standardized loadings: (* = significant at 1% level)

	f1	f2	f3	unique.var	communalities
A1		.*	*	0.902	0.098
A2		0.607*		0.615	0.385
A3		0.715*		0.483	0.517
A4	*	0.433*	.*	0.767	0.233
A5		0.681*	.*	0.496	0.504
C1	0.621*		*	0.604	0.396
C2	0.648*			0.575	0.425
C3	0.474*		.*	0.751	0.249
C4	-0.614*		0.371*	0.493	0.507
C5	-0.444*	*	0.406*	0.608	0.392
E1	*	-0.572*		0.685	0.315
E2		-0.619*	.*	0.510	0.490
E3	.*	0.615*		0.548	0.452
E4	*	0.720*	.*	0.451	0.549
E5	0.340*	0.442*		0.611	0.389

N1			0.859*	0.265	0.735
N2	*		0.840*	0.290	0.710
N3			0.751*	0.438	0.562
N4		.*	0.625*	0.533	0.467
N5	*		0.538*	0.700	0.300
O1	0.400*	.*		0.746	0.254
O2	-0.373*	*	.*	0.849	0.151
O3	0.386*	0.331*	*	0.678	0.322
O4	.*		.*	0.864	0.136
O5	-0.393*			0.841	0.159

	f2	f3	f1	total
Sum of sq (obliq) loadings	3.814	3.322	2.561	9.697
Proportion of total	0.393	0.343	0.264	1.000
Proportion var	0.153	0.133	0.102	0.388
Cumulative var	0.153	0.285	0.388	0.388

Factor correlations: (* = significant at 1% level)

	f1	f2	f3
f1	1.000		
f2	0.256*	1.000	
f3	-0.003	-0.143*	1.000

Number of factors: 4

Standardized loadings: (* = significant at 1% level)

	f1	f2	f3	f4	unique.var	communalities
A1		.*	*		0.904	0.096
A2	.*	0.592*			0.592	0.408
A3	.*	0.702*			0.470	0.530
A4	.*	0.423*		.*	0.670	0.330
A5		0.672*	.*		0.495	0.505
C1	0.563*			.*	0.593	0.407
C2	0.679*		*	.*	0.495	0.505
C3	0.592*				0.644	0.356
C4	-0.657*		.*	.*	0.447	0.553
C5	-0.569*	*	0.307*		0.524	0.476
E1	.*	-0.571*		*	0.675	0.325
E2		-0.616*	.*	.*	0.506	0.494
E3		0.626*		0.326*	0.506	0.494
E4		0.715*	.*		0.442	0.558

E5	.*	0.441*	.*	0.610	0.390
N1		0.866*	*	0.256	0.744
N2		0.843*		0.291	0.709
N3		0.751*		0.435	0.565
N4	.*	.*	0.606*	0.526	0.474
N5		0.567*	.*	0.651	0.349
O1		.*	0.571*	0.607	0.393
O2			.* -0.523*	0.689	0.311
O3		0.358*	0.648*	0.457	0.543
O4			.* 0.373*	0.792	0.208
O5			.* -0.580*	0.657	0.343

	f2	f3	f1	f4	total
Sum of sq (obliq) loadings	3.747	3.187	2.240	1.893	11.066
Proportion of total	0.339	0.288	0.202	0.171	1.000
Proportion var	0.150	0.127	0.090	0.076	0.443
Cumulative var	0.150	0.277	0.367	0.443	0.443

Factor correlations: (* = significant at 1% level)

	f1	f2	f3	f4
f1	1.000			
f2	0.237*	1.000		
f3	-0.090*	-0.137*	1.000	
f4	0.096*	0.006	0.057	1.000

Number of factors: 5

Standardized loadings: (* = significant at 1% level)

	f1	f2	f3	f4	f5	unique.var	communalities
A1	-0.472*		.*			0.744	0.256
A2	0.466*	*	0.565*	.*		0.467	0.533
A3	0.440*		0.685*	*		0.383	0.617
A4	.*	.*	0.401*		.*	0.643	0.357
A5	0.346*		0.677*			0.459	0.541
C1		0.578*			.*	0.593	0.407
C2		0.687*		.*	.*	0.492	0.508
C3		0.592*				0.643	0.357
C4		-0.673*		.*	.*	0.449	0.551
C5		-0.585*	*	0.318*		0.505	0.495
E1	.*	.*	-0.604*	*		0.598	0.402
E2	.*		-0.653*	.*		0.391	0.609

E3			0.665*		.*	0.500	0.500
E4			0.743*		.*	0.391	0.609
E5	.*	.*	0.477*		.*	0.542	0.458
N1	-0.530*			0.787*	*	0.193	0.807
N2	-0.471*			0.760*		0.284	0.716
N3	.*			0.756*		0.418	0.582
N4		.*	.*	0.655*	*	0.441	0.559
N5				0.603*	.*	0.601	0.399
O1			.*		0.545*	0.609	0.391
O2		*		.*	-0.518*	0.689	0.311
O3			0.409*		0.610*	0.461	0.539
O4	.*			0.301*	0.435*	0.656	0.344
O5				*	-0.584*	0.643	0.357

	f3	f4	f2	f5	f1	total
Sum of sq (obliq) loadings	3.832	3.012	2.224	1.795	1.344	12.207
Proportion of total	0.314	0.247	0.182	0.147	0.110	1.000
Proportion var	0.153	0.120	0.089	0.072	0.054	0.488
Cumulative var	0.153	0.274	0.363	0.435	0.488	0.488

Factor correlations: (* = significant at 1% level)

	f1	f2	f3	f4	f5
f1	1.000				
f2	0.078	1.000			
f3	-0.100	0.293*	1.000		
f4	0.113	-0.050	-0.130*	1.000	
f5	0.018	0.011	0.008	0.025	1.000

Number of factors: 6

Standardized loadings: (* = significant at 1% level)

	f1	f2	f3	f4	f5	f6	unique.var
A1	-0.607*	.*			.*		0.605
A2	0.648*		0.336*	.*			0.401
A3	0.504*		0.539*		*		0.392
A4	0.314*	.*	.*			.*	0.646
A5	.*		0.627*	.*			0.432
C1		0.598*				.*	0.583
C2		0.733*			.*		0.444
C3	*	0.586*				.*	0.648
C4	.*	-0.649*	.*	.*	0.392*		0.348

C5	-0.552*			0.337*	.*	0.515
E1	.*	-0.548*	-0.382*			0.553
E2		-0.621*	.*	.*		0.380
E3		0.728*		.*	.*	0.457
E4		0.759*	.*		.*	0.362
E5	.*	0.478*	.*		.*	0.544
N1			0.595*	0.756*	*	0.198
N2			0.578*	0.700*		0.258
N3			.*	0.749*		0.412
N4	.*	*	.*	0.630*	.*	0.436
N5	.*			* 0.610*	.*	0.595
O1		* 0.364*			0.534*	0.585
O2				0.313*	-0.514*	0.661
O3		0.467*			0.607*	0.459
O4	.*		.*	.*	0.456*	0.655
O5	.*			.*	-0.610*	0.575
communalities						
A1	0.395					
A2	0.599					
A3	0.608					
A4	0.354					
A5	0.568					
C1	0.417					
C2	0.556					
C3	0.352					
C4	0.652					
C5	0.485					
E1	0.447					
E2	0.620					
E3	0.543					
E4	0.638					
E5	0.456					
N1	0.802					
N2	0.742					
N3	0.588					
N4	0.564					
N5	0.405					
O1	0.415					
O2	0.339					
O3	0.541					
O4	0.345					
O5	0.425					

	f3	f5	f2	f6	f1	f4	total
Sum of sq (obliq) loadings	3.451	2.971	2.228	1.685	1.485	1.034	12.856
Proportion of total	0.268	0.231	0.173	0.131	0.116	0.080	1.000
Proportion var	0.138	0.119	0.089	0.067	0.059	0.041	0.514
Cumulative var	0.138	0.257	0.346	0.413	0.473	0.514	0.514

Factor correlations: (* = significant at 1% level)

	f1	f2	f3	f4	f5	f6
f1	1.000					
f2	0.177*	1.000				
f3	0.145*	0.268*	1.000			
f4	-0.092	0.034	-0.032	1.000		
f5	-0.015	-0.110	-0.289*	-0.121	1.000	
f6	0.022	0.164*	-0.089	0.121*	0.120*	1.000

With target rotation

```
# Target rotation (using the same matrix defined above)
efa5fac_fit <- efa(
  select(bfi, A1:O5),
  nfactors = 5,
  ordered = TRUE,
  rotation = "target",
  rotation.args = list(
    target = target_mat_bfi
  )
)
summary(efa5fac_fit)
```

This is lavaan 0.6-19 -- running exploratory factor analysis

Estimator	DWLS	
Rotation method	PST OBLIQUE	
Rotation algorithm (rstarts)	GPA (30)	
Standardized metric	TRUE	
Row weights	None	
	Used	Total
Number of observations	2436	2800

Fit measures:

```

          chisq  df pvalue   cfi rmsea
nfactors = 5 1856.317 185      0 0.914 0.066

```

Eigenvalues correlation matrix:

```

      ev1      ev2      ev3      ev4      ev5      ev6      ev7      ev8      ev9      ev10
5.725    2.960    2.294    1.964    1.638    1.050    0.797    0.767    0.669    0.637
      ev11     ev12     ev13     ev14     ev15     ev16     ev17     ev18     ev19     ev20
0.620    0.584    0.560    0.541    0.504    0.480    0.453    0.429    0.413    0.383
      ev21     ev22     ev23     ev24     ev25
0.354    0.335    0.328    0.310    0.203

```

Standardized loadings: (* = significant at 1% level)

	f1	f2	f3	f4	f5	unique.var	communalities
A1	-0.502*	*	.*	.*	*	0.744	0.256
A2	0.684*	*	*			0.467	0.533
A3	0.693*		.*		*	0.383	0.617
A4	0.467*	.*	.*		.*	0.643	0.357
A5	0.552*		.*	.*		0.459	0.541
C1		0.595*		*	.*	0.593	0.407
C2	*	0.705*	*	.*	*	0.492	0.508
C3	.*	0.606*	*	*	*	0.643	0.357
C4		-0.692*		.*		0.449	0.551
C5		-0.604*	.*	.*	.*	0.505	0.495
E1	*	.*	-0.636*			0.598	0.402
E2	*		-0.710*	.*		0.391	0.609
E3	.*		0.497*	*	.*	0.500	0.500
E4	.*		0.648*	*	.*	0.391	0.609
E5		.*	0.488*	.*	.*	0.542	0.458
N1	.*		.*	0.896*	*	0.193	0.807
N2	.*		.*	0.848*		0.284	0.716
N3	*	*		0.749*		0.418	0.582
N4	*	.*	-0.341*	0.555*	.*	0.441	0.559
N5	.*		.*	0.548*	.*	0.601	0.399
O1		*	.*		0.550*	0.609	0.391
O2	.*	*		.*	-0.508*	0.689	0.311
O3	*		.*		0.630*	0.461	0.539
O4	.*		.*	.*	0.474*	0.656	0.344
O5	*			.*	-0.591*	0.643	0.357

```

          f4      f3      f2      f1      f5  total
Sum of sq (obliq) loadings 3.029 2.650 2.368 2.293 1.868 12.207

```

Proportion of total	0.248	0.217	0.194	0.188	0.153	1.000
Proportion var	0.121	0.106	0.095	0.092	0.075	0.488
Cumulative var	0.121	0.227	0.322	0.414	0.488	0.488

Factor correlations: (* = significant at 1% level)

	f1	f2	f3	f4	f5
f1	1.000				
f2	0.216*	1.000			
f3	0.318*	0.284*	1.000		
f4	-0.087*	-0.206*	-0.213*	1.000	
f5	0.163*	0.204*	0.148*	0.020	1.000

APA Table

Openness and Agreeableness

```
fa_oblimin$loadings |>
  unclass() |>
  as.data.frame() |>
  tibble::rownames_to_column() |>
  flextable::flextable() |>
  bold(i = ~ abs(PA1) >= .30, j = ~ PA1) |>
  bold(i = ~ abs(PA2) >= .30, j = ~ PA2) |>
  set_formatter(
    PA1 = rmlead0,
    PA2 = rmlead0
  ) |>
  set_header_labels(values = c("item", "1", "2")) |>
  add_header_row(values = c("", "Factor loading"), colwidths = c(1, 2)) |>
  align(i = 1, align = "center", part = "header")
```

item	Factor loading	
	1	2
O1	.05	.51
O2	.14	-.49
O3	.11	.62
O4	-.02	.30

item	Factor loading	
	1	2
O5	.07	-.56
A1	-.35	-.04
A2	.67	.01
A3	.75	.00
A4	.51	-.10
A5	.62	.05

25-item from IPIP

```
fa_target_bfi$loadings |>
  unclass() |>
  as.data.frame() |>
  tibble::rownames_to_column() |>
  flextable::flextable() |>
  bold(i = ~ abs(PA1) >= .30, j = ~ PA1) |>
  bold(i = ~ abs(PA2) >= .30, j = ~ PA2) |>
  bold(i = ~ abs(PA3) >= .30, j = ~ PA3) |>
  bold(i = ~ abs(PA4) >= .30, j = ~ PA4) |>
  bold(i = ~ abs(PA5) >= .30, j = ~ PA5) |>
  set_formatter(
    PA1 = rmlead0,
    PA2 = rmlead0,
    PA3 = rmlead0,
    PA4 = rmlead0,
    PA5 = rmlead0
  ) |>
  set_header_labels(values = c("item", "1", "2", "3", "4", "5")) |>
  add_header_row(values = c("", "Factor loading"), colwidths = c(1, 5)) |>
  align(i = 1, align = "center", part = "header")
```

item	Factor loading				
	1	2	3	4	5
A1	.15	.18	.07	-.51	-.09

item	Factor loading				
	1	2	3	4	5
A2	.03	.05	.07	.69	.03
A3	.02	.18	.02	.69	.04
A4	-.02	.10	.21	.47	-.18
A5	-.11	.28	-.00	.55	.04
C1	.07	-.02	.59	-.00	.17
C2	.18	-.05	.71	.06	.06
C3	.04	-.08	.61	.10	-.07
C4	.17	.03	-.69	.01	-.03
C5	.21	-.10	-.60	.01	.12
E1	-.01	-.62	.13	-.07	-.02
E2	.17	-.71	-.01	-.04	.01
E3	.07	.51	-.02	.22	.27
E4	-.05	.64	-.00	.27	-.13
E5	.13	.50	.28	.01	.17
N1	.86	.21	.01	-.19	-.08
N2	.81	.14	.02	-.16	.00
N3	.77	.00	-.03	.02	.01
N4	.56	-.33	-.13	.08	.13
N5	.56	-.17	-.00	.20	-.15
O1	-.00	.18	.06	-.00	.55
O2	.20	.02	-.08	.15	-.50
O3	.02	.27	-.01	.05	.63
O4	.19	-.28	-.04	.19	.48
O5	.11	.05	-.04	.04	-.59